

Dealing With Vine's Issues

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- 1 Vine Copulas
 - Vine Copula Definition
 - Vine Copula Example
- 2 Lattice Selection
 - Principle
 - Example
- 3 Numerical Application

General Copula Definition

Definition: Copulas

- Let $X = [X_1, X_2, \dots, X_n]$ be a vector of random variables:
 - of distribution function F ,
 - of joint density f ,
 - of marginal distribution functions F_1, F_2, \dots, F_n ,
 - of marginal densities f_1, f_2, \dots, f_n .
- Sklar (1959)'s theorem insures the existence of two functions c and C such that:

$$F(x) = C(F_1(x_1), F_2(x_2), \dots, F_n(x_n))$$

$$f(x) = \prod_{i=1}^n f_i(x_i) \cdot c(F_1(x_1), F_2(x_2), \dots, F_n(x_n))$$

Vine Copula Definition

Definition: Vine Copulas

A vine copula is a n -variate parametric copula built by decomposing the multivariate density in a product of bivariate copulas.

- Origins: Joe (1997) and Bedford and Cooke (2002, 2001)
- Applications: Aas et al. (2009), Berg and Aas (2009), Czado et al. (2009) and Guégan and Maugis (2009).

Decomposition Example In Dimension 3

$$\begin{aligned}f(x_1, x_2, x_3) &= f(x_1) \cdot f(x_2) \cdot f(x_3) \\ &\cdot c_{12}(F_1(x_1), F_2(x_2)) \cdot c_{13}(F_1(x_1), F_3(x_3)) \\ &\cdot c_{23|1}(\partial_1 C_{12}(F_1(x_1), F_2(x_2)), \partial_1 C_{13}(F_1(x_1), F_2(x_2)))\end{aligned}$$

Trivariate Vine Copula

$$\begin{aligned}C_{123} &= C_{12}(F_1(x_1), F_2(x_2)) \cdot c_{13}(F_1(x_1), F_3(x_3)) \\ &\cdot c_{23|1}(\partial_1 C_{12}(F_1(x_1), F_2(x_2)), \partial_1 C_{13}(F_1(x_1), F_2(x_2))) \\ C_{123} &= C_{12} \cdot C_{13} \cdot C_{23|1}\end{aligned}$$

Issues:

- Number of vine copulas:

$$\frac{n!}{2} \prod_{i=1}^{n-3} i!$$

- Copula Selection: Chen et al. (2004) and Chen and Fan (2006) and Kiefer and Choi (2006).

Lattice Selection:

Gabriel (1969), Andersson (1988), Edwards and Havranek (1987) and Antoch and Hanousek (2000)

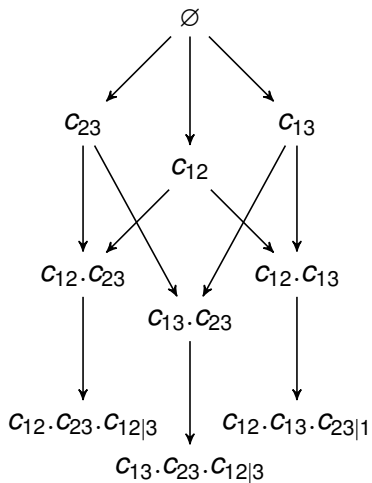
Specification & Generalization: Lattice structure

Among models, some can be considered as specifications of others, and some as generalizations of others.

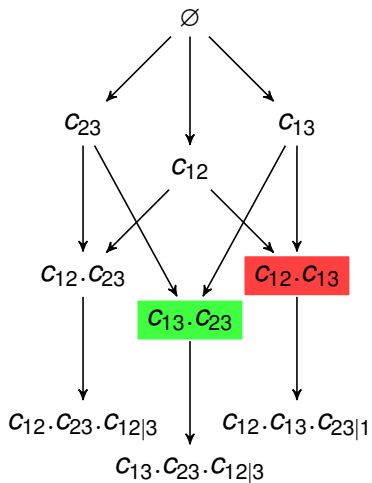
Coherence Principle:

Specifications of a false model are false, and generalizations of valid models are valid.

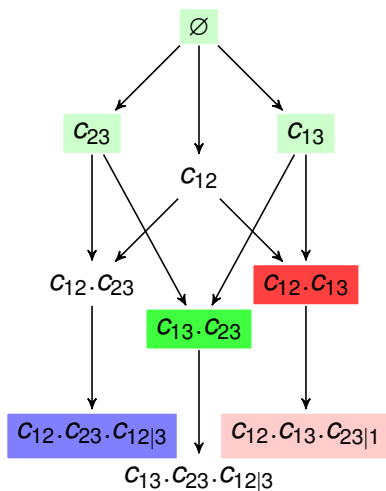
Example



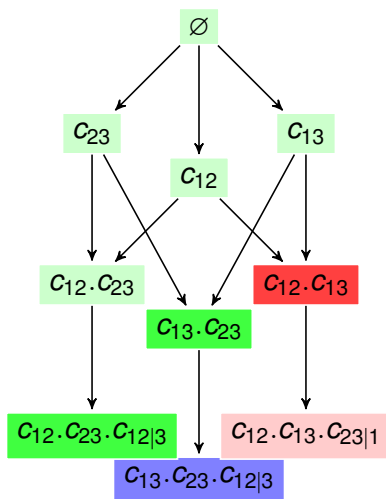
Example



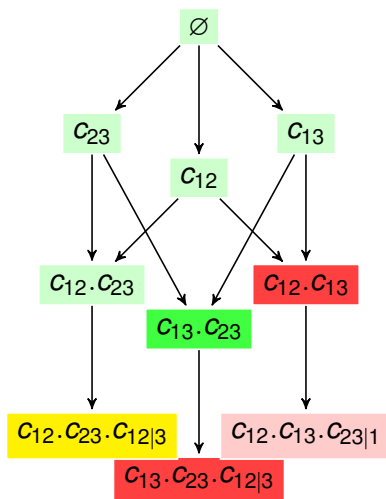
Example



Example

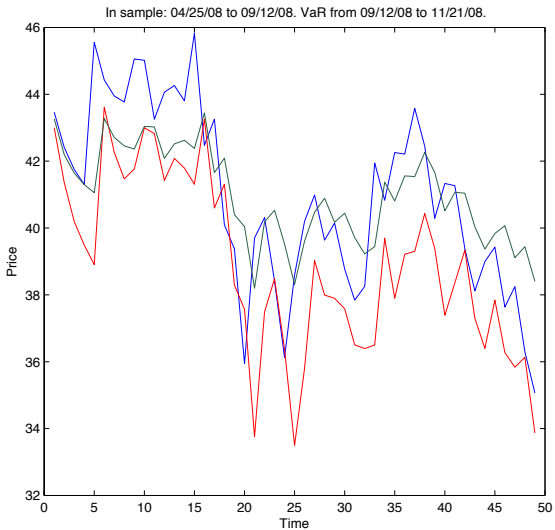


Example



Procedure:

- Five main assets composing the *CAC40* from 25/04/08 to 21/11/08 (Datastream).
- In-sample: 04/25/08 to 09/12/08.
Out-Sample: 09/12/08 to 11/21/08.
- Fully automated estimation and search.
 - Univariate process: *GARCH* using *AIC*.
 - Multivariate estimation: Maximum Likelihood.
 - Test: Modified Anderson Darling (Chen et al., 2004).



Summary

- Vine copulas are parametric copulas built using bivariate copulas.
- Vine copulas can be selected using a lattice based search algorithm.
- By completely modeling the multivariate distribution, Vine copulas permits to build better risk measures.

Issues & Prospects

Issues:

Heavily Computational.

Prospects:

Better bivariate copulas.

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Annex 1

$$f(x_1, x_2, x_3) = f(x_1, x_2) \cdot \frac{f(x_2, x_3|x_1)}{f(x_2|x_1)}$$

$$f(x_1, x_2, x_3) = f(x_1) \cdot f(x_2) \cdot c_{12}(F_1(x_1), F_2(x_2)) \\ \cdot \frac{f(x_2|x_1) \cdot f(x_3|x_1)}{f(x_2|x_1)} \cdot c_{23|1}(F_{2|1}(x_2|x_1), F_{3|1}(x_3|x_1))$$

$$f(x_1, x_2, x_3) = f(x_1) \cdot f(x_2) \cdot c_{12}(F_1(x_1), F_2(x_2)) \\ \cdot \frac{f(x_1, x_3)}{f(x_1)} \cdot c_{23|1}(\partial_1 c_{12}(F_1(x_1), F_2(x_2)), \partial_1 c_{13}(F_1(x_1), F_2(x_2)))$$

$$f(x_1, x_2, x_3) = f(x_1) \cdot f(x_2) \cdot f(x_3) \\ \cdot c_{12}(F_1(x_1), F_2(x_2)) \cdot c_{13}(F_1(x_1), F_3(x_3)) \\ \cdot c_{23|1}(\partial_1 c_{12}(F_1(x_1), F_2(x_2)), \partial_1 c_{13}(F_1(x_1), F_2(x_2)))$$

Annex 2 I

The test is the following: we are testing the null hypothesis H_0 against the alternative hypothesis H_1 :

$$H_0 : Pr(C(U_1, \dots, U_n) = C_0(U_1, \dots, U_n)) = 1$$

$$H_1 : Pr(C(U_1, \dots, U_n) = C_0(U_1, \dots, U_n)) < 1$$

where C is the true copula and C_0 the estimated copula. We define then $\{Z_i\}_{i < n}$ and W as:

$$Z_i = F(U_i | U_1, \dots, U_{i-1}) \text{ and } W = \sum_1^n [\Phi^{-1}(Z_i)]^2,$$

where Φ is the standard normal distribution function. Then W follows a χ_d^2 distribution; our test will be based on this result. For this purpose we will use this univariate boundary kernel $K_h(x, y)$

$$K_h(x, y) = \begin{cases} k(\frac{x-y}{h}) / \int_{-\frac{x}{h}}^1 k(u) du & \text{if } x \in [0, h) \\ k(\frac{x-y}{h}) & \text{if } x \in [h, 1-h] \\ k(\frac{x-y}{h}) / \int_{-\frac{x}{h}}^1 k(u) du & \text{if } x \in (1-h, 1] \end{cases}$$

Annex 2 II

in which $k(\cdot)$ is the quartic kernel: $k(u) = \frac{15}{16}(1 - u^2)^2 \mathbf{1}_{|u| < 1}$. Then we define g_W as a kernel estimation of its density function, where h is fixed using the rule of thumb ($h = \sqrt{\text{Var}(W)} n^{-1/5}$):

$$g_W(\omega) = \frac{1}{n * h} \sum_t^T K_h(\omega, F_{\chi_d^2}(W_t))$$

The alternative test is based on:

$$J_n = \int_0^1 [g_W(\omega) - 1]^2 d\omega$$

Then under classical hypothesis we have that:

$$\text{Stat}_n = \frac{(nh^{1/2} J_n - r_n)}{\sigma} \rightarrow N(0, 1) \text{ in distribution.}$$

Where:

$$r_n = h^{1/2} \left[(h^{-1} - 2) \int_{-1}^1 k^2(\omega) d\omega + 2 \int_0^1 \int_{-1}^z k_z^2(y) dy dz \right]$$

$$\sigma^2 = 2 \int_{-1}^1 \left[\int_{-1}^1 k(u+v)k(v) dv \right]^2 du$$

$$k_z(y) = k(y) / \int_{-1}^z k(u) du$$

For further details on the test, including Monte-Carlo estimation of the power of the test, see Chen et al. (2004).